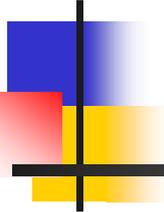


Statistical Mechanical Approach to CDMA Multiuser Detection Algorithm



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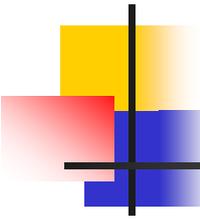
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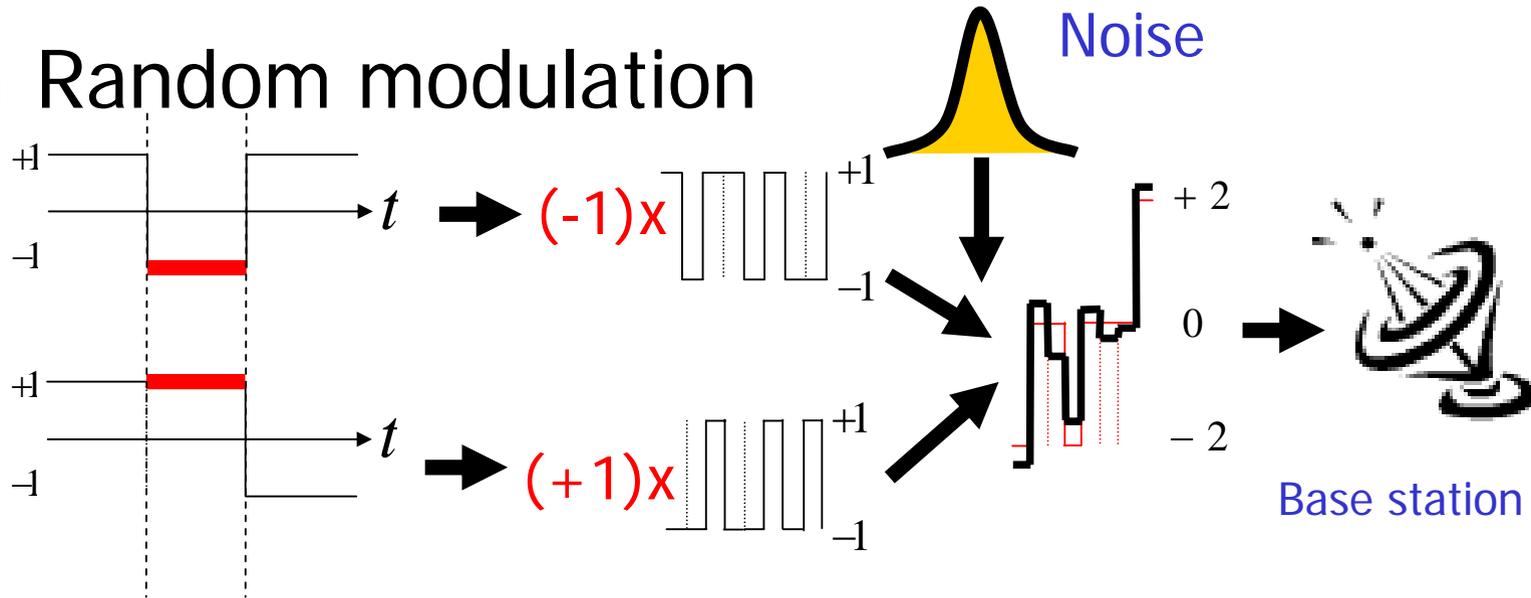


Outline

- CDMA system and multi-user detection
- Graphical expression & belief propagation (BP)
- Statistical-mechanical approach
- Extension to survey propagation (SP)
- Summary

DS/BPSK CDMA system

■ Random modulation

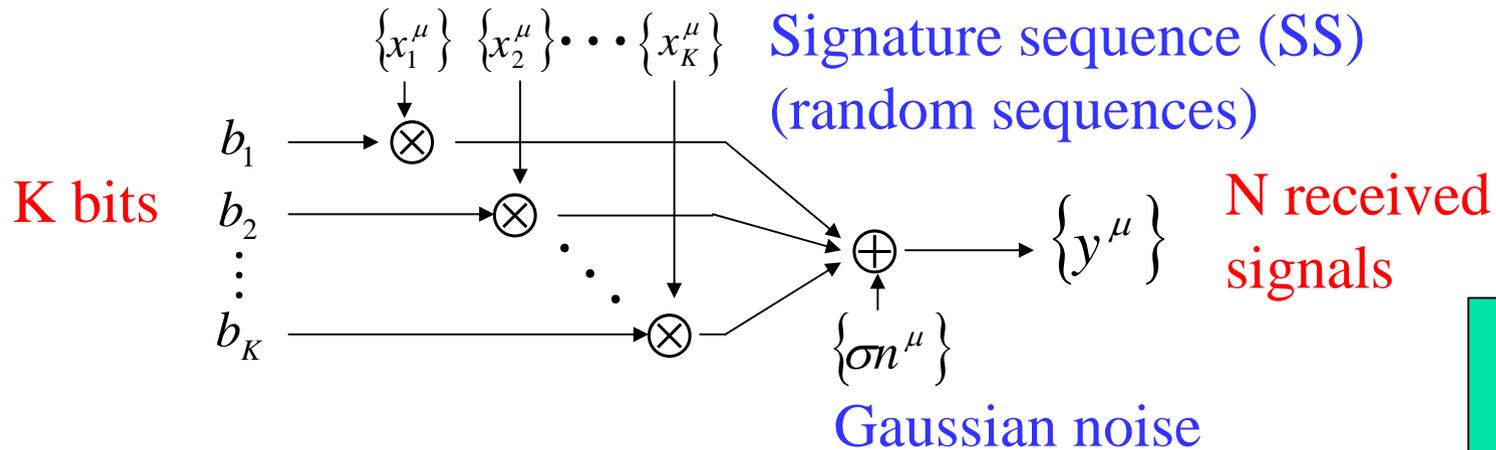


Modulate each signal with
N random binary sequences
(W-CDMA: up to $N=256$)

$\mathbf{X} = (x_1, x_2, \dots, x_8)$

Mathematical model

- Multi-user (**many-to-one**) communication



$$y^\mu = \frac{1}{\sqrt{N}} \sum_{i=1}^K x_i^\mu b_i + \sigma n^\mu$$

Problem: Detect \mathbf{b} from \mathbf{y} .

What is required for the user detection

- Real-time communication
 - Detection must be quickly performable
 - Low error rate is preferred
- As b_i are **binary variables**, development of detection schemes satisfying these requirements is non-trivial

$$y^\mu = \frac{1}{\sqrt{N}} \sum_{i=1}^K x_i^\mu b_i + \sigma n^\mu$$

Linear equation:
difficult to solve
for discrete variables

A useful property

- Random signature sequences

$$x_i^\mu : \text{i.i.d. from } P(x) = \frac{1}{2}\delta(x + 1) + \frac{1}{2}\delta(x - 1)$$



$$\frac{1}{N} \mathbf{x}_i \cdot \mathbf{x}_j = \frac{1}{N} \sum_{\mu=1}^N x_i^\mu x_j^\mu = \begin{cases} 1, & i = j \\ O(N^{-1/2}) & i \neq j \end{cases}$$

- Random SSs are nearly orthogonal to each other

User i's SS

$$\mathbf{x}_i = \{x_i^\mu\}$$

User j's SS

$$\mathbf{x}_j = \{x_j^\mu\}$$



Single-user detection

- Convenient method to detect b_i (in use)
 - Operate \mathbf{x}_i to the received signals $\{y^\mu\}$

$$h_i = \frac{1}{\sqrt{N}} \sum_{\mu=1}^N x_i^\mu y^\mu = b_i + \sum_{j \neq k} \frac{1}{N} \sum_{\mu=1}^N x_i^\mu x_j^\mu b_j + \frac{1}{\sqrt{N}} \sum_{\mu=1}^n x_i^\mu \sigma n^\mu$$

Signal

Cross-talk noise

Channel noise



$$\hat{b}_i^{SD} = \text{sign}(h_i)$$

small: $O(N^{-1/2})$

- Good news
 - Quickly performable using only the focused user's SS
 - In use in standard systems
- Bad news
 - High bit error rate (BER) when # of users is large.

#Better scheme is demanded for high-perform. comm.

Optimal detection

- Bayesian approach

Uniform Prior

$$P(\mathbf{b}) = 2^{-K}$$

Gaussian channel

$$P(y^\mu | \mathbf{x}^\mu, \mathbf{b}) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^\mu - \frac{\mathbf{x}^\mu \cdot \mathbf{b}}{\sqrt{N}})^2}{2\sigma^2}}$$

$$P(\mathbf{b} | D^N) = \frac{P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})}{\sum_{\mathbf{w}} P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})} \propto \exp \left[-\frac{1}{2\sigma^2} \sum_{\mu=1}^N \left(y^\mu - \frac{\mathbf{x}^\mu \cdot \mathbf{b}}{\sqrt{N}} \right)^2 \right]$$

- Optimal detection (Max. Posterior Marginals: MPM)

$$\hat{b}_i = \operatorname{argmax}_{b_i} \left\{ \sum_{\mathbf{b} \setminus b_i} P(\mathbf{b} | D^N) \right\}$$

#Minimizes BER $\operatorname{Prob}(\hat{b}_i \neq b_i)$

Computational difficulty

- Unfortunately, performing the optimal detection is computationally difficult!

$$P(\mathbf{b}|D^N) = \frac{P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})}{\sum_{\mathbf{b}} P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})}$$

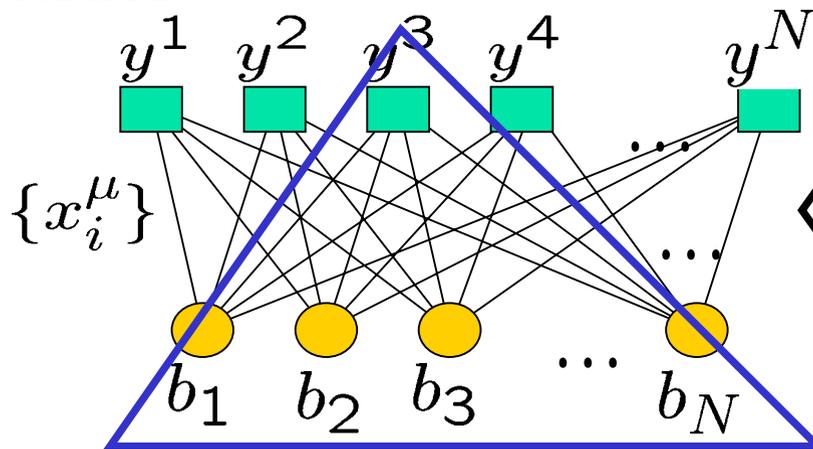
$O(2^N)$ summations!

- Development of good *approx. algorithms* is necessary.

Graphical representation

- In order to answer this, we introduce a **factor (bipartite) graph** expression of the posterior dist.

Data nodes



Posterior dist.

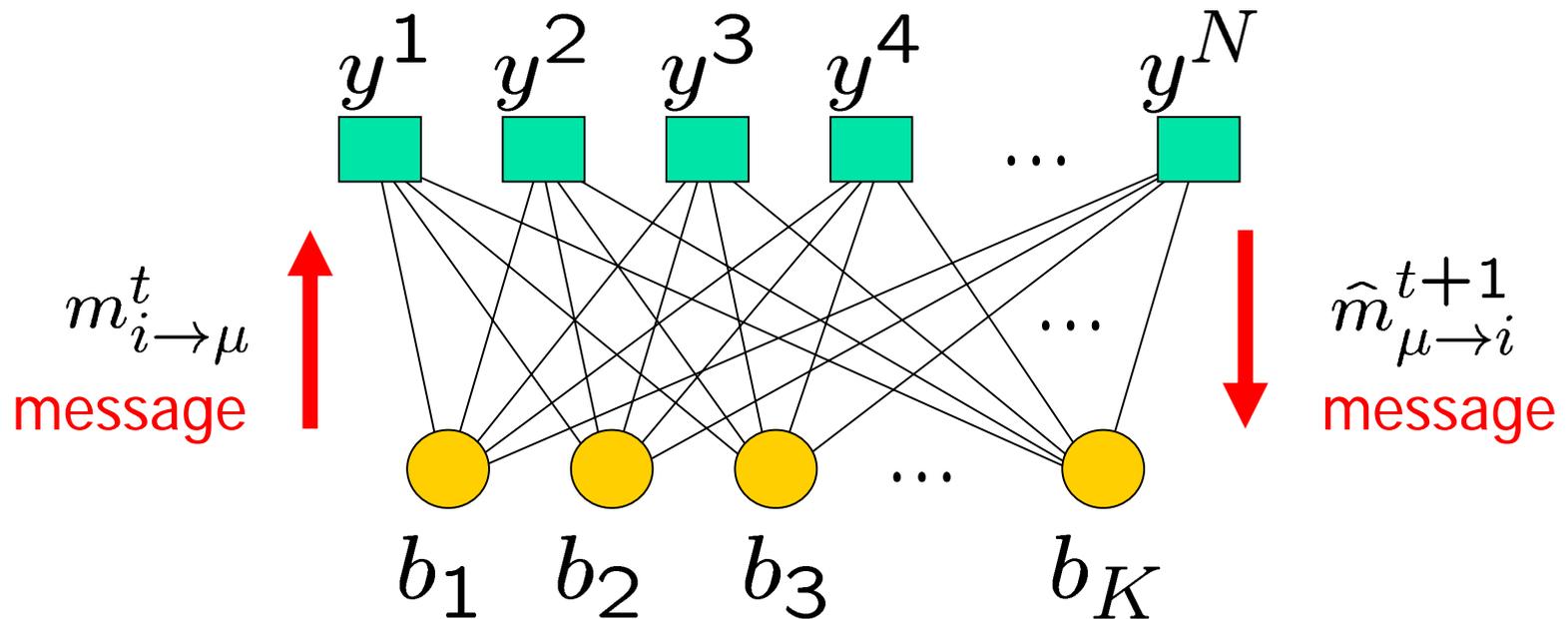
$$P(\mathbf{b} | D^N) = \frac{P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})}{\sum_{\mathbf{b}} P(\mathbf{b}) \prod_{\mu=1}^N P(y^\mu | \mathbf{x}^\mu, \mathbf{b})}$$

$$\Leftrightarrow P(y^3 | \mathbf{x}^3, \mathbf{b})$$

Variable nodes

Belief Propagation (BP)

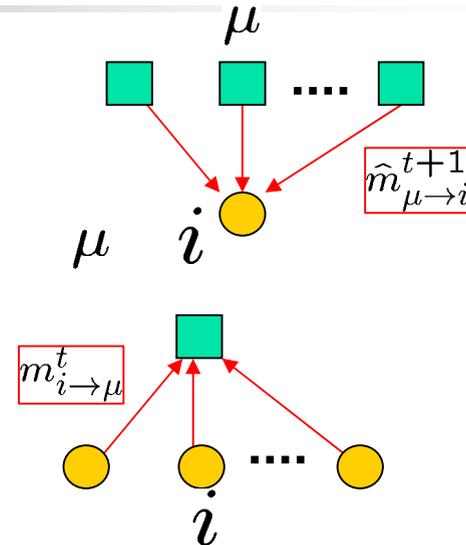
- Pearl (1987), MacKay (1995)
- Iteratively passing *messages* between the two types of nodes.



BP-based detection

- More specifically,

$$\begin{cases} \hat{m}_{\mu \rightarrow i}^{t+1} = \frac{\sum_{\mathbf{b}} b_i P(y^\mu | \mathbf{x}^\mu, \mathbf{b}) \prod_{j \neq i} \left(\frac{1 + m_{j \rightarrow \mu}^t b_j}{2} \right)}{\sum_{\mathbf{b}} P(y^\mu | \mathbf{x}^\mu, \mathbf{b}) \prod_{j \neq i} \left(\frac{1 + m_{j \rightarrow \mu}^t b_j}{2} \right)} \\ m_{i \rightarrow \mu}^t = \tanh \left[\sum_{\nu \neq \mu} \tanh^{-1}(\hat{m}_{\nu \rightarrow i}^t) \right] \end{cases}$$



- Bayes detection

$$m_i^t = \langle b_i \rangle^t = \tanh \left[\sum_{\mu=1}^N \tanh^{-1}(\hat{m}_{\mu \rightarrow i}^t) \right] \text{ : Posterior mean}$$

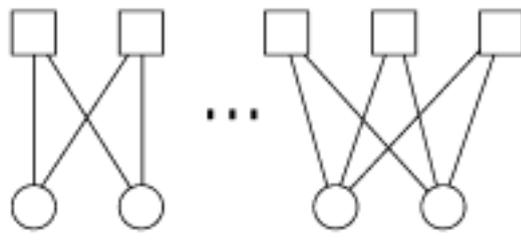
$$\hat{b}_i^t = \operatorname{argmax}_{b_i} \left\{ \sum_{\mathbf{b} \setminus b_i} P^t(\mathbf{b} | D^N) \right\} = \operatorname{sign}(m_i^t)$$

- As information of all users is required, this is among **multi-user detection algorithms**.

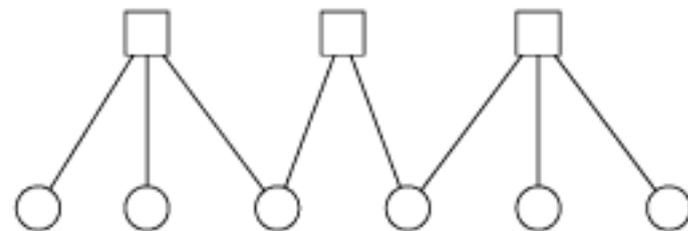
Theorem (Pearl 1987)

- BP provides the exact average for **loop-free graphs** after messages propagate once over the graph.
 - BP can be also employed in **loopy graphs** as an **approximation** algorithm (**loopy BP**)

Loops in graphs

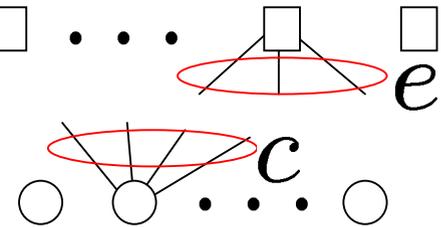


Loop-free graph



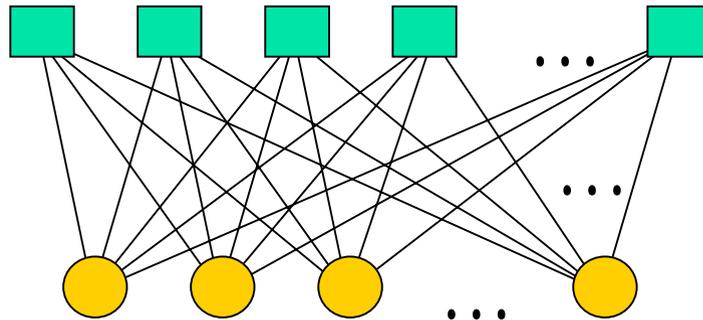
Intuitive speculation about loopy BP

- As BP provides the exact result for loop free graphs, the lower the density of short loops in the graph is, the better performance will be gained.
- Cf) Random sparse graphs:
 - Loop lengths $O(\ln N)$.
 - Clique sizes $O(1)$.
- This speculation is empirically confirmed in several applications such as **LDPC codes**.



Pessimistic perspective

- CDMA system = Complete bipartite graph



Two bad news

- *Many Short loops*
 - *Highly dense*
- Seeing these, you may feel that the current BP-based approach is not promising...
 - However, techniques from S.M. can develop a nearly optimal algorithm based on BP, making use of the denseness of the graph appropriately when the network size is large.

Gaussian approximation

- **Key property**

$$P(y^\mu | \mathbf{x}^\mu, \mathbf{b}) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^\mu - \frac{\mathbf{b} \cdot \mathbf{x}^\mu}{\sqrt{N}})^2}{2\sigma^2}} \quad \text{depends on } \mathbf{b}$$

only through $\frac{\mathbf{b} \cdot \mathbf{x}^\mu}{\sqrt{N}}$.

- When $b \sim \prod_j \left(\frac{1+m_{j \rightarrow \mu}^t b_j}{2} \right)^{\sqrt{N}}$ and N is large, $\frac{\mathbf{b} \cdot \mathbf{x}^\mu}{\sqrt{N}}$ can be handled as a Gaussian variable (Mezard 1989, Opper and Winther 1996).

- This provides the following Gaussian approximation

$$\sum_{\mathbf{b}} P(y^\mu | \mathbf{x}^\mu, \mathbf{b}) \prod_{j \neq i} \left(\frac{1+m_{j \rightarrow \mu}^t b_j}{2} \right) \simeq \int \frac{d\Delta^\mu e^{-\frac{(\Delta^\mu - \langle \Delta^\mu \rangle_\mu^t)^2}{2V_\mu^t}}}{\sqrt{2\pi V_\mu^t}} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y^\mu - \Delta^\mu)^2}{2\sigma^2}}$$

$O(2^N)$ compt



$O(1)$ compt.

Drastic reduction!

Stat. mech. algorithm

- Algorithm of $O(K^2N)$ computations/update

$$\left\{ \begin{array}{l} a_{\mu \rightarrow i}^{t+1} = \frac{1}{\sigma^2 + \beta(1 - Q^t)} \left(y^\mu - \sum_{j \neq i} \frac{x_j^\mu m_{j \rightarrow \mu}^t}{\sqrt{N}} \right) \\ m_{i \rightarrow \mu}^t = \tanh \left[\sum_{\nu \neq \mu} \frac{x_i^\nu}{\sqrt{N}} a_{\nu \rightarrow i}^t \right] \end{array} \right.$$

$$m_i^t = \tanh \left[\sum_{\mu=1}^N \frac{x_i^\mu}{\sqrt{N}} a_{\mu \rightarrow i}^t \right]$$

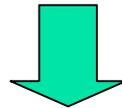
$$\left(\begin{array}{l} \beta = \frac{K}{N} \sim O(1) \\ Q^t \equiv \frac{1}{N} \sum_{i=1}^N (m_i^t)^2 \end{array} \right)$$

Further reduction of compt. cost

- Taylor expansion

$$m_{i \rightarrow \mu}^t \simeq m_i^t - \left(1 - (m_i^t)^2\right) \underbrace{\tilde{m}_{\mu \rightarrow i}^t}_{\text{small}}$$

- This makes it possible to express the BP update using only the singly-indexed variables



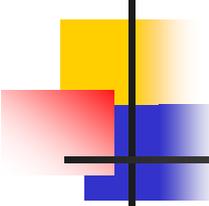
Further reduction to $O(NK)$ computations/update

Faster BP-based algorithm

- Algorithm of $O(KN)$ computations/update

$$\left\{ \begin{array}{l} a_{\mu}^{t+1} = \frac{1}{\sigma^2 + \beta(1 - Q^t)} \left(y^{\mu} - \sum_{i=1}^K \frac{x_i^{\mu}}{\sqrt{N}} m_i^t + \beta(1 - Q^t) a_{\mu}^t \right) \\ m_i^t = \tanh \left(\sum_{\mu=1}^N \frac{x_i^{\mu}}{\sqrt{N}} a_{\mu}^t + \frac{m_i^{t-1}}{\sigma^2 + \beta(1 - Q^{t-1})} \right) \end{array} \right.$$

#Expressed by only singly-indexed variables



Remark (I)

- Computational cost is similar to that of a conventional algorithm
- Multi-stage detection (Varanasi and Aazhang 1991)

$$\left\{ \begin{array}{l} a_{\mu}^{t+1} = y^{\mu} - \sum_{i=1}^K \frac{x_i^{\mu}}{\sqrt{N}} m_i^t \\ m_i^t = \text{sign} \left(\sum_{\mu=1}^N \frac{x_i^{\mu}}{\sqrt{N}} a_{\mu} + m_i^{t-1} \right) \end{array} \right.$$

Remark (II)

- The fixed point of the BP-based algorithm is characterized by the **TAP equation** of this system.

$$m_i = \tanh \left[\frac{1}{\sigma^2} \left(h_i - \sum_{j \neq i} J_{ij} m_j \right) - \frac{\beta(1-Q)m_i}{\sigma^2(\sigma^2 + \beta(1-Q))} \right]$$

$$\left(h_i = \frac{1}{\sqrt{N}} \sum_{\mu=1}^N x_i^\mu y^\mu, \quad J_{ij} = \frac{1}{N} \sum_{\mu=1}^N x_i^\mu x_j^\mu, \quad Q = \frac{1}{K} \sum_{i=1}^K m_i^2 \right)$$

Remark (III)

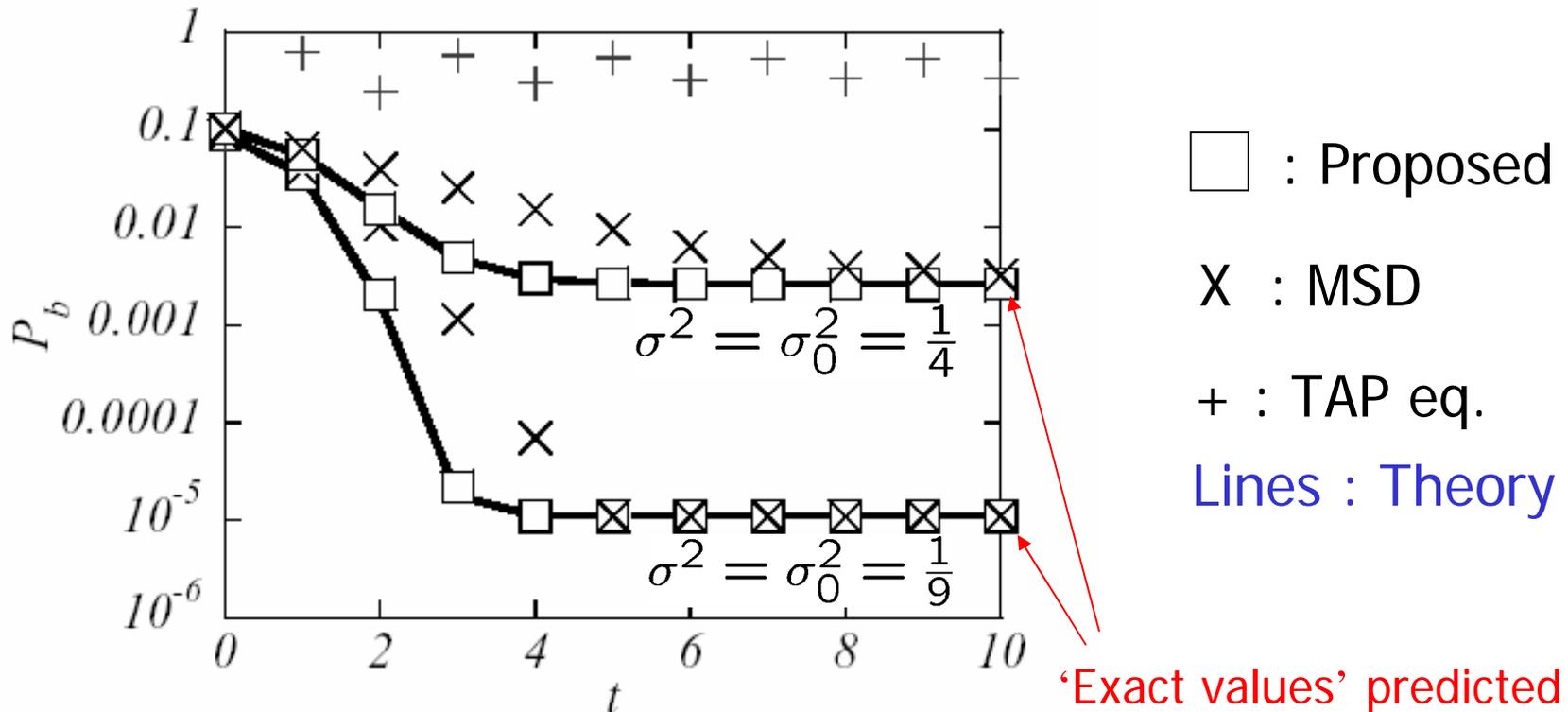
- Macroscopic dynamics (**Density Evolution**) can be well captured by the iteration of SP eq. of the **replica symmetric (RS) analysis** (Tanaka 2002).

$$\begin{cases} E^{t+1} = \frac{1}{\sigma^2 + \beta(1-Q^t)}, & F^{t+1} = \frac{\beta(1-2M^t + Q^t) + \sigma_0^2}{[\sigma^2 + \beta(1-Q^t)]^2} \\ M^t = \int Dz \tanh(\sqrt{F^t} z + E^t), & Q^t = \int Dz \tanh^2(\sqrt{F^t} z + E^t) \end{cases}$$

- This implies that the developed approx. algorithm practically converges to the ‘exact’ solution when Nishimori’s condition $\sigma^2 = \sigma_0^2$ holds, for which no RSB is empirically expected.

Method comparison

- Condition: $K = 1000$ $N = 2000$



Remark (IV)

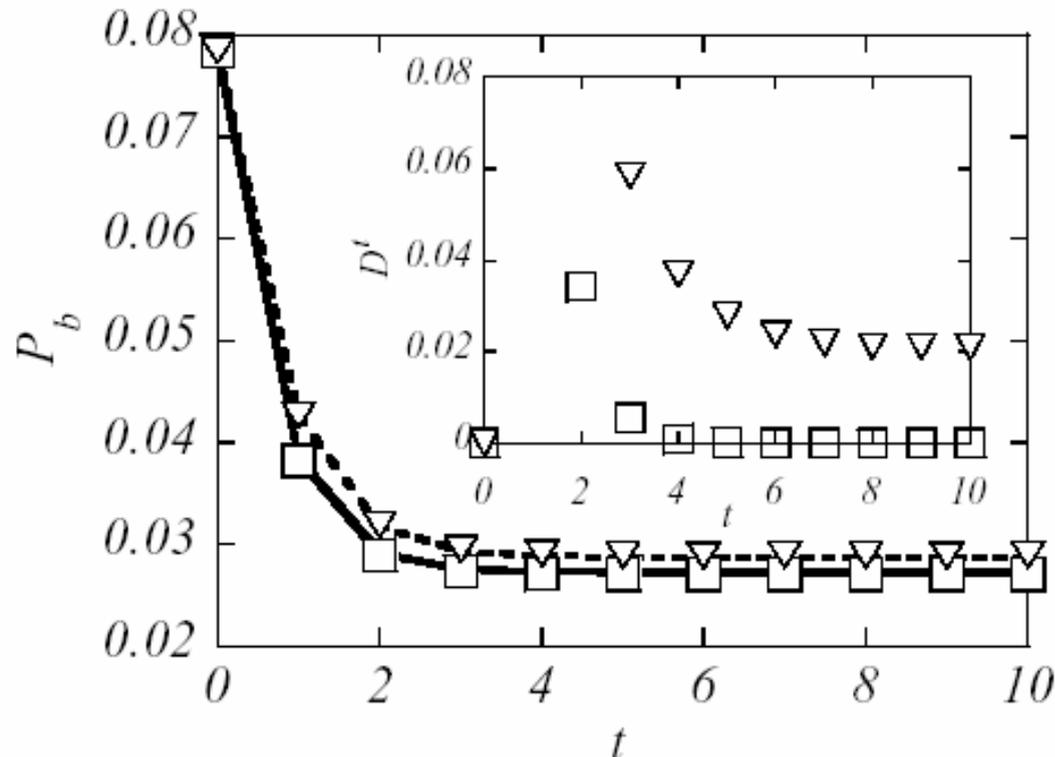
- Microscopic (**dynamical**) instability condition of the fixed point becomes identical to that of the **AT instability** of the (**equilibrium**) replica analysis
 - This can occur when the assumed noise parameter σ^2 is sufficiently smaller than the true one σ_0^2 .

Microscopic Instability Condition(=AT Condition)

$$\frac{1}{[\sigma^2 + \beta(1 - Q)]^2} \times \frac{1}{N} \sum_{i=1}^K (1 - m_i^2)^2 > 1$$

Microscopic instability

$$K = 500, N = 2000, \sigma_0^2 = 0.25$$



$$\nabla: \sigma^2 = 0.01$$

$$\square: \sigma^2 = \sigma_0^2 = 0.25$$

$$D^t = \frac{1}{K} \sum_{i=1}^K (m_i^t - m_i^{t-1})^2$$

Measures Micro. Instability

Extension to Survey Prop.

- Link to the AT instability motivates us to develop an algorithm based on the survey propagation (SP), which can describe RSB phase.
- **SP (Mezard-Parisi-Zecchina 2002)**
 - Distributions of messages (surveys)
 - Introduction of RSB parameter \mathcal{X}
 - Correspondence to the 1RSB solution
- In the current system, the central limit theorem makes it possible to express SP compact.

Remarks on SP-based algorithm

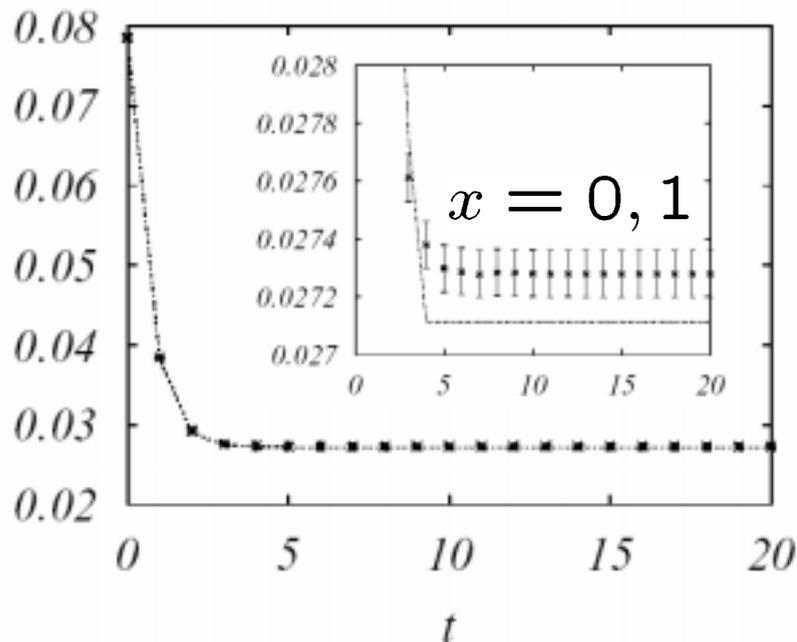
- Under tree approximation, the macroscopic dynamics is described by natural iteration of 1RSB SP eqs.
 - The microscopic instability does not correspond to the AT instability of the 1RSB solution.
- Robust for mismatch of noise parameter σ^2
 - AT stable: as good as BP independently of RSB param. \mathcal{X}
 - AT unstable: can be better than BP
- Tuning \mathcal{X} by the free energy maximization principle is not necessarily optimal for reducing BER.
- $x = 1$ and $x = 0$ provide different performance when BP is unstable although both of these param. choices are reduced to the RS solution in the replica analysis.

$x = 1$ vs $x = 0$

$$\beta = K/N = 500/2000$$

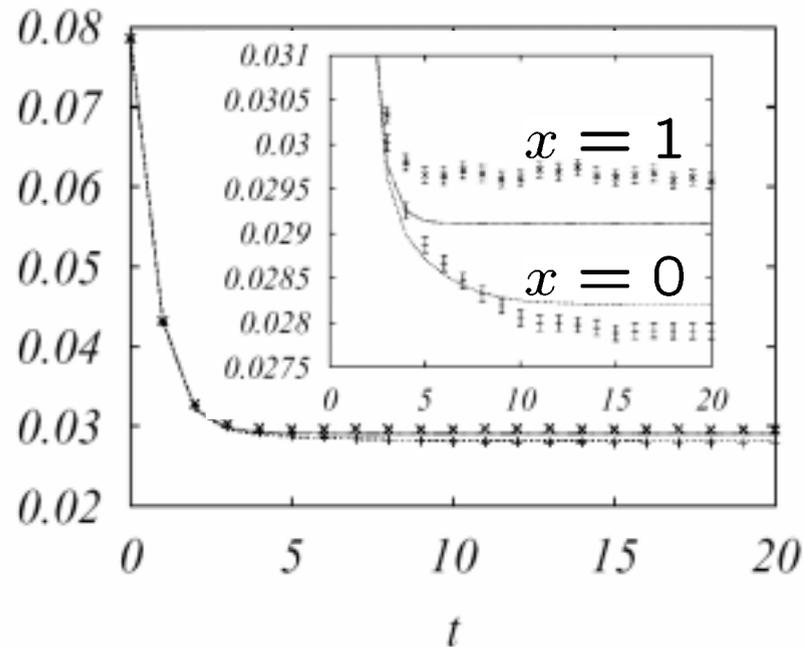
RS regime

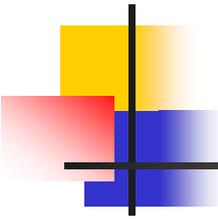
$$(\sigma_0^2, \sigma^2) = (0.25, 0.25)$$



RSB regime

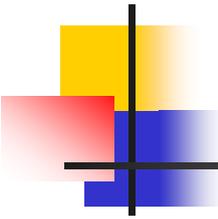
$$(\sigma_0^2, \sigma^2) = (0.25, 0.01)$$





Summary

- Approx. algorithm for the Bayes detection in CDMA system based on BP and SM.
 - Quicker convergence.
 - Compt. cost is not significantly increased.
 - Excellent consistency with the replica theory.
 - Converges to the optimal solution in the thermodynamic limit if the assumed hyper parameters are correct.
- Extension to SP
 - SP can serve as a robust algorithm for mismatch of hyper parameters.



References

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